

Research

Psychiatric morbidity and cervical cancer screening: a retrospective population-based case-cohort study

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Abstract

Background: Cervical cancer screening reduces disease-specific mortality. This study aimed to estimate whether bipolar disorder or schizophrenia is associated with disparities in cervical cancer screening rates.

Methods: This was a retrospective population-based matched case—cohort study of community-dwelling women aged 19–69 in Ontario using linked health administrative databases. We used odds ratios (ORs), hazards ratios and rate ratios (RRs) adjusted for demographic characteristics and relevant comorbidities to compare cervical cancer screening outcomes between women with a diagnosis of bipolar disorder or schizophrenia to women without that history matched on key demographic characteristics, between 2003 and 2015.

Results: In total, 1 245 457 women were identified for inclusion in the analyses, 119 948 with a diagnosis of bipolar disorder or schizophrenia, and 1 125 509 without. Over a median follow-up duration of 12.5 years, women with the exposure were 36% less likely to be screened (OR 0.64, 95% confidence interval [CI] 0.64–0.65) than those without, and they took longer to undergo screening (median 18.98 mo v. 16.63 mo; χ^2 = 3718.2, p < 0.001). They were also screened less frequently (median 6.16 yr v. 4.69 yr per screen; RR 0.85, 95% CI 0.84–0.85). These effects were consistent after we excluded the 86 475 women (6.9%) with suspected major depressive disorder, and they were larger for the 59 141 women (4.7%) not attached to a family physician.

Interpretation: Women with bipolar disorder or schizophrenia were less likely to undergo cervical cancer screening, their screening was delayed, and they were screened at a lower rate compared to women without this psychiatric history. This practice gap suggests a need to further address barriers to screening, including access to a family physician, among women with bipolar disorder or schizophrenia.

n Canadian women, mortality related to cervical cancer decreased by 60% between 1977 and 2006, largely attributable to provincial screening programs that facilitate detection and treatment of precancerous lesions. ^{1,2} The 2005 Ontario Cancer Screening Program practice guideline recommended that cervical cytology screening be initiated within 3 years of first vaginal sexual activity. ³ Screening was recommended every 3 years following 3 consecutive negative results of annual Papanicolaou tests. ³ This guideline was updated in 2012, specifying that sexually active women aged more than 21 should be screened every 3 years and removing the criteria for negative results of 3 consecutive annual tests. ⁴ Adherence has been moderate, with 65% of eligible Ontario women receiving a screen between 2009 and 2011. ⁵

Women with severe psychiatric conditions have a significantly higher rate of death from cervical cancer than those without these conditions despite comparable disease incidence.⁶ Previous reports show that women with common and severe psychiatric conditions, including bipolar disorder and schizophrenia, face barriers to screening uptake;^{7,8} however, estimates have varied.^{1,9-11} Different selection criteria,

classification methods, durations of follow-up or other factors may affect generalizability to the general population. 9,11

Further population-based estimates are required to accurately quantify potential disparities in screening practices between women with and without psychiatric morbidity, and to promote the development of targeted interventions.^{6,9,11} In the present study, we aimed to compare the likelihood and frequency of cervical cancer screening in community-dwelling women with and without a history of bipolar disorder or schizophrenia from the general population in Ontario.

Competing interests: Jessica Colby-Milley (formerly of Sunnybrook Research Institute) is now an employee of F. Hoffmann-La Roche (effective July 29, 2019). She receives full-time salary support from F. Hoffmann-La Roche. This role is unrelated to the present work. Her contributions to this manuscript were made before she commenced her role at F. Hoffmann-La Roche. No other competing interests were declared.

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Methods

Study design and data sources

This was a retrospective matched case–cohort study using linked provincial administrative databases from Ontario, which has a population of about 14 million. ¹² All data were obtained from ICES, and women were identified across databases by means of unique encoded patient identifiers. The data set from this study is held securely in coded form at ICES. Access may be granted to those who meet prespecified criteria for confidential access, available at www.ices.on.ca/DAS.

We used data from the Canadian Institute for Health Information Discharge Abstract Database, National Ambulatory Care Reporting System and Ontario Health Insurance Plan (OHIP) database, which capture universally available government-funded coverage for all hospital and emergency department services, physician visits and diagnostic tests, combined with the Ontario Mental Health Reporting System Metadata, to identify bipolar disorder and schizophrenia diagnoses, using the International Classification of Diseases, 9th revision (ICD-9) and the International Statistical Classification of Diseases and Related Health Problems, 10th revision (ICD-10). 13,14 We obtained demographic information from the Ontario Registered Persons Database, and data on prescription history (for those aged ≥ 65 yr) from the Ontario Drug Benefit Claims database. We used the Ontario Breast Screening Program database for outcome assessment to identify all cervical cancer screening procedures conducted during the study period.

These databases are subject to routine assessments of data quality.¹⁵ The Canadian Institute for Health Information provides facility-level reports for the Discharge Abstract Database,¹⁶ National Ambulatory Care Reporting System¹⁷ and Ontario Mental Health Reporting System,¹⁸ which contain complete or nearly complete data for Ontario. These reports identify data quality issues to improve data accuracy and comparability.

Study population and exposure

We identified and followed community-dwelling women aged 19-69 with a diagnosis of bipolar disorder or schizophrenia between 2003 and 2012 for the primary outcome of cervical cancer screening until the age of 70 (when cessation of screening was recommended in Ontario guidelines throughout the study period¹⁹), death, loss of OHIP eligibility or termination of follow-up on Dec. 31, 2015. Cohort entry at the age of 19 is consistent with the prevailing Ontario guideline¹⁹ and with other cohort studies that have defined entry between age 18 and 21.9,10,20,21 Exclusion criteria included an invalid OHIP number, and a history of cervical cancer or hysterectomy before July 1 of the year of cohort entry. The Ontario Health Insurance Plan is the taxpayer-funded government-operated health insurance plan that provides universal coverage of preventive and emergency physician and hospital services for people whose principal residence is in Ontario and who are present in Ontario for more than half of the days in a given year. Coverages can be lost if these eligibility criteria cease to be fulfilled.

The exposure was a prior diagnosis of bipolar disorder or schizophrenia, defined with the following algorithm: a discharge diagnosis of bipolar disorder or schizophrenia from the Discharge Abstract Database or Ontario Mental Health Reporting System using ICD-9 or ICD-10 diagnostic codes^{13,14} (Appendix 1, Supplemental Table S1, available at www.cmajopen.ca/content/8/1/E134/suppl/DC1) within a 10-year period before entry into the study cohort; an OHIP diagnostic billing code of bipolar disorder or schizophrenia (Appendix 1, Supplemental Table S1), with accompanying specialty code for psychiatry (19), within a 12-month period within 10 years before cohort entry; or 2 OHIP diagnostic billing codes of bipolar disorder or schizophrenia within a 12-month period within 10 years before cohort entry.

We matched women with bipolar disorder or schizophrenia within each geographic Local Health Integration Network at a ratio of up to 10:1 to community-dwelling women from the general Ontario population (n = 1 125 509) on age (± 2 yr), Charlson Comorbidity Index score, income quintile, urban or rural residence, attachment to a family physician and Adjusted Diagnostic Group classification using greedy nearest-neighbour matching.²² Local Health Integration Networks are nonprofit community organizations defined by 14 separate geographic regions in Ontario. They oversee, regulate and fund public health care services and facilities, including hospitals, long-term care and mental health services, to improve patient access and the patient experience. Women without the exposure were matched only once with women who had the exposure; thus, owing to imbalances in the size of our subgroups, some exposed subjects could not be matched at an exact 10:1 ratio.

Outcome

The outcome was a documented screen for cervical cancer, defined with OHIP procedure code G365, E430 or G394.

Statistical analysis

We generated descriptive statistics to characterize the study cohort with respect to demographic and clinical variables. We compared baseline characteristics between women with and without the exposure using standardized differences and χ^2 tests. We determined incidence rates of cervical cancer screening per 3 person-years to assess adherence to Ontario guideline for cervical cancer screening intervals.¹⁹

We used conditional logistic regression to estimate adjusted odds ratios (ORs) with 95% confidence intervals (CIs) of the likelihood of cervical cancer screening. We used Poisson regression to compare adjusted rate ratios (RRs) with 95% CIs of screening between exposed and unexposed women, and Cox proportional hazards regression to estimate adjusted hazard ratios (HRs) with 95% CIs for time to screening. We estimated these parameters using data within 3 years of cohort entry and for all available follow-up data. We used a log-rank test to compare time to first screen. Models were adjusted for certain potential clinical confounders that may be more likely to occur in women with psychiatric morbidity than in those without, or that may





affect overall medical burden or propensity to receive cervical screening (Appendix 1, Supplemental Table S1).

To probe the influence of accessibility of health care resources, we performed subgroup analyses based on whether the woman was attached to a family physician. In this analysis, physician attachment was not used as a factor for matching.

Depressive disorders are common psychiatric comorbidities,^{23,24} and they have been independently associated with reduced use of health care services.^{25–28} However, inconsistent use of diagnostic codes for major depressive disorder can result in misclassification;²⁹ therefore, we did not include suspected major depressive disorder in the exposure. In a sensitivity analysis, we excluded women with suspected major depressive disorder (ICD-9 code F296.2–296.36, ICD-10 code F33.x) to determine whether suspected major depressive disorder affected observed screening rates and associations.

We performed all statistical analyses using SAS, Version 14.2 (SAS Institute).

Ethics approval

This study was approved by the Research Ethics Board of Sunnybrook Health Sciences Centre. The details of this study are reported in accordance with STROBE and RECORD guidelines.^{30,31}

Results

Table 1 shows baseline characteristics for the whole cohort of Ontario women aged 19–69 on cohort entry (n = 1 245 457), the 119 948 women with bipolar disorder or schizophrenia, and the 1 125 509 matched women without the exposure. Of the 119 948 women with bipolar disorder or schizophrenia, 98 368 (82.0%) were matched 10:1, and 21 580 (18.0%) were matched between 1:1 and 9:1. The demographic characteristics of women with bipolar disorder or schizophrenia were similar to those of women without these diagnoses, although the prevalence of vascular diseases, diabetes, dyslipidemia, respiratory diseases, arthritis, osteoporosis, obesity, pelvic inflammatory diseases and polycystic ovary syndrome was significantly higher among exposed v. unexposed women (Table 1).

Likelihood, rate and time of screening

Within the first 3 years of cohort entry, 719 240 women (57.7%) underwent cervical cancer screening, 60 436 (50.4%) of those with schizophrenia or bipolar disorder, and 658 804 (58.5%) of those without. The adjusted OR for having received 1 screen within the first 3 years was 0.71 (95% CI 0.70–0.72) (Table 2). The rate of screening within the first 3 years was 0.27 (95% CI 0.27–0.28) screens per 3 person-years among women with bipolar disorder or schizophrenia, and 0.36 (95% CI 0.36–0.37) among those without (RR 0.80, 95% CI 0.80–0.81) (Table 3 and Table 4). For comparison, the independent effects of other chronic diseases are presented in Table 4.

Throughout the entire follow-up period (median 12.5 yr), 958 249 women (76.9%) received at least 1 screen (84 228

[70.2%] in the exposure group and 874 021 [77.7%] in the group without exposure). The adjusted OR for having received at least 1 screen for women with schizophrenia or bipolar disorder relative to women without the exposure was 0.64 (95% CI 0.64–0.65) (Table 2). The rate of screening was 0.49 (95% CI 0.48–0.49) screens per 3 person-years (i.e., on average, screening occurred every 6.16 yr) among women with bipolar disorder or schizophrenia, and 0.64 (95% CI 0.64–0.64) (i.e., screening every 4.69 yr) among those without (RR 0.85, 95% CI 0.84–0.85) (Table 3 and Table 4).

Among women who were screened at least once throughout the follow-up period, the rate of screening was lower for those with bipolar disorder or schizophrenia than for those without (HR 0.80, 95% CI 0.80–0.80) (Table 5, Figure 1A). Specifically, women with the exposure were screened a median of 18.98 months after cohort entry, whereas those without the exposure were screened after a median of 16.63 months (log-rank test χ^2 = 3718.22, p < 0.001). The median length of time to either first screening after cohort entry or termination of follow-up was significantly longer for women with a history of schizophrenia or bipolar disorder (30.70 mo) than for those without (23.40 mo) (p < 0.001).

Physician attachment

Among the 1 186 316 women (95.3%) attached to a family physician, the adjusted OR for receiving 1 screen in the first 3 years was 0.72 (95% CI 0.71–0.73); throughout the entire follow-up period, it was 0.65 (95% CI 0.64-0.66) (Appendix 1, Supplemental Table S2). In contrast, women not attached to a physician had a lower OR for screening over the first 3 years (0.55, 95% CI 0.51-0.59) and over the entire followup period (0.55, 95% CI 0.52–0.58) (Appendix 1, Supplemental Table S3). The RR of screening over all follow-up associated with bipolar disorder or schizophrenia was 0.85 (95% CI 0.85–0.86) for women attached to a family physician and 0.67 (95% CI 0.64-0.70) for those not attached to a family physician (Appendix 1, Supplemental Table S4). The adjusted HR of screening was 0.81 (95% CI 0.80-0.81) for attached women and 0.64 (95% CI 0.61-0.67) for unattached women (Appendix 1, Supplemental Table S5).

Among women who were attached to a family physician, the rate of screening was 0.521 screens (95% CI 0.517–0.524) per 3 person-years (i.e., screening every 5.76 yr) for those with the exposure and 0.681 screens (95% CI 0.680–0.683) (i.e., screening every 4.40 yr) for those without. Among women not attached to a family physician, the corresponding rates were 0.139 (95% CI 0.133–0.145) (i.e., screening every 21.63 yr) and 0.210 (95% CI 0.207–0.213) (i.e., screening every 14.29 yr).

Sensitivity analysis among women without major depressive disorder comorbidity

A total of 86 475 women (6.9%) in the overall cohort had a history of comorbid depressive illness (8516 [7.1%] of those exposed and 78 229 [7.0%] of those unexposed) (Appendix 1, Supplemental Table S6). After these women were excluded, the likelihood of screening remained lower for women with



Table 1: Baseline demographic and health characteristics of women in the general Ontario population aged 19–69 years between July 2003 and July 2012, comparing those with bipolar disorder or schizophrenia to those without

	Group; no. (%) of women*			_
	Overall n = 1 245 457	Exposure n = 119 948	Matched n = 1 125 509	Standardized difference†
Age, yr, median (IQR)	42 (29–52)	42 (30–52)	42 (29–52)	0.02
Income quintile				
1 (lowest)	298 115 (23.9)	29 896 (24.9)	268 219 (23.8)	0.03
2	265 181 (21.3)	25 548 (21.3)	239 633 (21.3)	< 0.001
3	236 935 (19.0)	22 443 (18.7)	214 492 (19.1)	0.01
4	224 190 (18.0)	21 186 (17.7)	203 004 (18.0)	0.01
5 (highest)	219 010 (17.6)	20 601 (17.2)	198 409 (17.6)	0.01
Missing	2026 (0.2)	274 (0.2)	1752 (0.2)	0.02
Residence				
Urban	1 131 150 (90.9)	108 389 (90.3)	1 022 761 (90.9)	0.01
Rural	113 879 (9.1)	11 496 (9.6)	102 383 (9.1)	0.02
Missing	428 (0.03)	63 (0.1)	365 (0.03)	0.01
Charlson Comorbidity Index score				
0	178 386 (14.3)	24 961 (20.8)	153 425 (13.6)	0.19
1	18 447 (1.5)	3451 (2.9)	14 996 (1.3)	0.11
2	10 749 (0.9)	1299 (1.1)	9450 (0.8)	0.02
≥3	8226 (0.7)	1002 (0.8)	7224 (0.6)	0.02
Missing	1 029 649 (82.7)	89 235 (74.4)	940 414 (83.6)	0.23
Adjusted Diagnostic Group score				
0	32 505 (2.6)	3011 (2.5)	29 494 (2.6)	0.01
1–3	137 524 (11.0)	12 762 (10.6)	124 762 (11.1)	0.01
4–6	318 091 (25.5)	29 369 (24.5)	288 722 (25.7)	0.03
7–9	371 691 (29.8)	34 265 (28.6)	337 426 (30.0)	0.03
≥ 10	385 646 (31.0)	40 541 (33.8)	345 105 (30.7)	0.07
Comorbidities				
Vascular diseases	15 082 (1.2)	1898 (1.6)	13 184 (1.2)	153.04
Hypertension	44 240 (3.6)	3720 (3.1)	40 520 (3.6)	78.72
Diabetes	18 251 (1.5)	2173 (1.8)	16 078 (1.4)	110.18
Dyslipidemia	4504 (0.4)	618 (0.5)	3886 (0.4)	86.90
Respiratory diseases	8390 (0.7)	1740 (1.4)	6650 (0.6)	1197.55
Arthritis	2547 (0.2)	302 (0.2)	2245 (0.2)	14.53
Crohn disease/ulcerative colitis	2135 (0.2)	198 (0.3)	1937 (0.2)	0.31
Osteoporosis	724 (0.1)	138 (0.1)	586 (0.05)	74.02
Obesity	4646 (0.4)	942 (0.8)	3704 (0.3)	607.13
Pelvic inflammatory diseases	870 (0.1)	115 (0.1)	755 (0.1)	12.87
Polycystic ovary syndrome	126 (0.01)	27 (0.02)	99 (0.01)	20.15
Attached to family physician	, ,	, ,	, ,	
Yes	1 186 316 (95.3)	113 722 (94.8)	1 072 594 (95.3)	0.01
No	59 141 (4.7)	6226 (5.2)	52 915 (4.7)	0.01

Note: IQR = interquartile range. *Except where noted otherwise. $+\chi^2$ test for comorbidities. For other characteristics, we calculated standardized differences through the difference in the mean between the 2 groups divided by an estimate of the standard deviation.



a psychiatric history than for those without a psychiatric history, both in the first 3 years (OR 0.72, 95% CI 0.71-0.73) and over all available follow-up (OR 0.65, 95% CI 0.64–0.66). Similarly, the adjusted HR for time to screening remained significantly lower (0.81, 95% CI 0.80-0.82) (Table 5, Figure 1B).

Table 2: Odds ratios from conditional logistic regression of screening likelihood within the first 3 years and over all available follow-up (2003-2015) for women with schizophrenia or bipolar disorder relative to women without the exposure

	Adjusted OR (95% CI)*		
Variable	Within first 3 yr	All follow-up	
Schizophrenia/bipolar disorder	0.71 (0.70–0.72)	0.64 (0.64–0.65)	
Vascular diseases	0.47 (0.45-0.49)	0.45 (0.43-0.46)	
Hypertension	0.92 (0.90-0.94)	0.92 (0.90-0.95)	
Diabetes	0.73 (0.71–0.76)	0.73 (0.70-0.76)	
Dyslipidemia	0.86 (0.80-0.93)	0.91 (0.84–0.98)	
Respiratory diseases	0.59 (0.56-0.62)	0.52 (0.49-0.55)	
Arthritis	0.76 (0.70-0.84)	0.86 (0.78-0.95)	
Crohn disease/ ulcerative colitis	0.75 (0.68–0.83)	0.77 (0.68–0.86)	
Osteoporosis	0.52 (0.43-0.63)	0.43 (0.35–0.51)	
Obesity	0.57 (0.53-0.61)	0.54 (0.50-0.58)	
Pelvic inflammatory diseases	0.66 (0.56–0.77)	0.58 (0.49–0.69)	
Polycystic ovary syndrome	0.93 (0.63–1.38)	1.04 (0.60–1.80)	

Note: CI = confidence interval OB = odds ratio *Adjusted for comorbidities included in table

Interpretation

For women with bipolar disorder or schizophrenia, we report a 36% reduced likelihood of receiving at least 1 cervical cancer screen and a 15% lower RR of screening over the 12-year follow-up period, which translates into an overall adherence to triannual screening recommendations that was about 24% lower than that for women without this psychiatric history. The time to first screen after cohort entry was longer among women with these psychiatric conditions than among matched, unexposed cases, which further confirms a reduced propensity for cervical cancer screening. The magnitude of these disparities was greater for women not attached to a family physician.

Women with and without bipolar disorder or schizophrenia were matched for important demographic characteristics, and the models were adjusted for vascular and other clinical comorbidities that may directly affect complexity of care or screening rates. Although most comorbid chronic conditions studied were more prevalent among women with bipolar disorder or schizophrenia than among women without these conditions, their presence did not account for the disparity in screening. In the adjusted models, the independent effect of bipolar disorder or schizophrenia on the rate of screening was similar to those of respiratory diseases, arthritis, osteoporosis and obesity, smaller than that of vascular diseases, and larger than those of hypertension, diabetes and dyslipidemia.

The findings confirm and quantify a gap in use of preventive health services between Ontario women with and without severe psychiatric disorders, and they contribute to a growing body of literature that identifies mental health conditions as barriers to the uptake of cervical cancer screening. 1,9,11,20,32,33 Notably, a study of women with and without schizophrenia in Manitoba identified a comparable magnitude of disparity (30%) in the likelihood of receiving cervical screening.1 A conflicting investigation¹⁰ showed a higher frequency of cervical cancer screening among women with psychosis, bipolar disorder or

Table 3: Incidence of screening and average screening frequency over the first 3 years of cohort entry and over all available follow-up (2003-2015) for the overall cohort and by exposure status

Group; analytical period	No. of screens	Person-years	Incidence per 3 person-years (95% CI)	Average years per screen
Overall				
3 yr	719 240	6 071 289	0.355 (0.355–0.356)	8.45
All follow-up	958 249	4 615 795	0.623 (0.622-0.624)	4.82
Schizophrenia/bipolar disorder				
3 yr	60 436	661 669	0.274 (0.272-0.276)	10.95
All follow-up	84 288	519 348	0.487 (0.483-0.490)	6.16
No schizophrenia/bipolar disorder				
3 yr	658 804	5 409 619	0.365 (0.364-0.366)	8.22
All follow-up	874 021	4 096 447	0.640 (0.639-0.641)	4.69

Table 4: Rate ratios from Poisson regression of rates of cervical cancer screening within the first 3 years and over all available follow-up (2003–2015) in women with schizophrenia or bipolar disorder relative to women without the exposure

	Adjusted RR (95% CI)*		
Variable	Within 3 yr	All follow-up	
Schizophrenia/bipolar disorder	0.80 (0.80-0.81)	0.85 (0.84–0.85)	
Vascular diseases	0.66 (0.65–0.67)	0.75 (0.74–0.76)	
Hypertension	0.95 (0.94–0.95)	0.96 (0.96–0.97)	
Diabetes	0.87 (0.86-0.88)	0.91 (0.90-0.92)	
Dyslipidemia	0.89 (0.86–0.92)	0.91 (0.89-0.93)	
Respiratory diseases	0.81 (0.80-0.83)	0.86 (0.85-0.87)	
Arthritis	0.82 (0.79-0.84)	0.87 (0.85-0.89)	
Crohn disease/ ulcerative colitis	0.94 (0.91–0.97)	0.96 (0.94–0.98)	
Osteoporosis	0.76 (0.70-0.82)	0.82 (0.78-0.87)	
Obesity	0.79 (0.77–0.81)	0.85 (0.83-0.86)	
Pelvic inflammatory diseases	0.91 (0.86–0.97)	0.96 (0.92-0.99)	
Polycystic ovary syndrome	1.03 (0.91–1.16)	1.03 (0.97–1.10)	
Note: CI = confidence interva	I BB = rate ratio		

Note: CI = confidence interval, RR = rate ratio.
*Adjusted for comorbidities included in table.

Table 5: Hazard ratios from Cox proportional hazards regression for time to first cervical cancer screen in the complete cohort and in a subgroup with major depressive disorder excluded

	Adjusted HR (95% CI)*		
Variable	Complete cohort	Subgroup excluding major depressive disorder	
Schizophrenia/bipolar disorder	0.80 (0.80–0.81)	0.81 (0.80-0.82)	
Vascular diseases	0.61 (0.60-0.63)	0.62 (0.60-0.64)	
Hypertension	0.94 (0.93-0.96)	0.94 (0.93–0.95)	
Diabetes	0.82 (0.80-0.84)	0.82 (0.81-0.84)	
Dyslipidemia	0.88 (0.84-0.92)	0.87 (0.83-0.92)	
Respiratory diseases	0.73 (0.70-0.75)	0.72 (0.70-0.75)	
Arthritis	0.85 (0.80-0.90)	0.84 (0.79-0.90)	
Crohn disease/ ulcerative colitis	0.85 (0.80–0.90)	0.85 (0.80–0.90)	
Osteoporosis	0.63 (0.55-0.72)	0.63 (0.55–0.72)	
Obesity	0.68 (0.65–0.71)	0.69 (0.66–0.72)	
Pelvic inflammatory diseases	0.81 (0.73–0.89)	0.81 (0.73–0.89)	
Polycystic ovary syndrome	0.95 (0.76–1.18)	0.93 (0.73–1.16)	

Note: CI = confidence interval, HR = hazard ratio. *Adjusted for comorbidities included in table.

mania; however, those cross-sectional data were drawn from an independent insurer-paid model with exclusively low-income participants, which suggests that demographic characteristics and care model may affect screening rates.

Future directions for research include identifying and validating other actionable barriers that could be targeted to diminish the observed disparity. Prior studies have suggested that both patient and physician efforts to address acute psychiatric symptoms may supersede indications for preventive procedures. To women with psychosis-related delusions and hallucinations, the purpose of screening or treatment may be unclear, and fear or distrust toward primary care physicians and invasive procedures may result in avoidance of screening. The values and attitudes of physicians can also play a role, and stigma can translate into disparities in screening and other services. 33,36

Evidence suggests that distributing targeted letters of invitation, offering telephone counselling, addressing financial barriers and focusing on continuity of care may improve uptake, 32,37 but the effectiveness of these methods has yet to be investigated specifically in women with severe psychiatric disorders. 38 In the present study, lack of attachment to a family physician was associated with a much larger effect of psychiatric morbidity on screening rates, identifying a group at particularly high risk and highlighting the importance of access to primary care. Future studies might stratify analyses based on other sociodemographic factors to identify subgroups that may be at the highest risk and to identify specific barriers to address to mitigate underuse.

There may be a role for family physicians to audit their practices to identify their rate of cervical cancer screening in this defined population and to adopt quality-improvement plans to target this screening disparity when identified. Psychiatrists, social workers and mental health professionals might encourage patients to visit family physicians for screening and advocate to improve primary care access.

Limitations

The population-based nature, use of public health records and relatively long follow-up are notable strengths of the present work. We also acknowledge several limitations. Ontario residents have universal access to physician and hospital care under OHIP, which limits the generalizability of our findings to populations under other models of care; however, the results provide estimates that are less likely to reflect differences in insurance availability, unemployment and socioeconomic status. The current analysis matched on Local Health Integration Network to further account for geographic and, to some extent, socioeconomic barriers, but we were unable to control for several individual sociodemographic factors that have been linked to underuse of health care services, including household income, cultural differences and comorbid substance use disorders.³⁹ In addition, although the specific algorithm used to ascertain the exposure in this study was not validated previously, the method is supported by prior work.^{40–42}

We did not ascertain estimates for major depressive disorder alone; however, the observed differences associated with



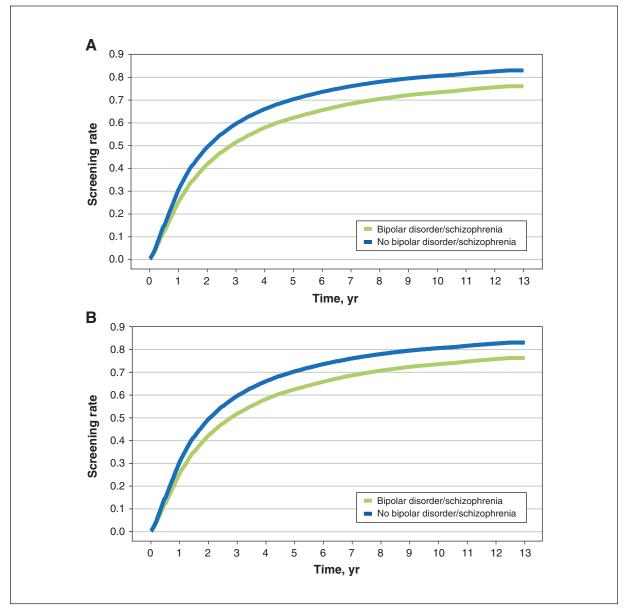


Figure 1: Cumulative incidence curve for cervical cancer screening over the follow-up period (2003-2015) for all women with and without bipolar disorder or schizophrenia (A) and excluding women with major depressive disorder (B). Log-rank test p < 0.001.

bipolar disorder or schizophrenia remained consistent in sensitivity analyses that excluded women with a diagnostic code for major depressive disorder.

Finally, as a general limitation of studies using public health records, it is possible that not all people with the exposure were identified; however, potential misclassification of some exposed women as unexposed would result in conservative estimates that underestimate the discrepancies identified.

Conclusion

The likelihood of cervical cancer screening was 36% lower among Ontario women with bipolar disorder or schizophrenia than among carefully matched women without these conditions, and this disparity was larger among those who were

not attached to a physician. Because screening is critical for the early detection and treatment of cervical cancer, this disparity is likely to result in increased morbidity and mortality, which warrants the attention of patients, their care providers and other health care professionals.

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