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Sex differences in the outcomes of peripheral arterial disease: a population-based cohort study

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Abstract

Background: The role of sex in the outcomes of patients with peripheral arterial disease (PAD) has been poorly studied. We sought to investigate differences in the long-term adverse cardiovascular and limb outcomes between men and women with PAD.

Methods: We conducted a population-based cohort study with up to 7 years of follow-up using linked administrative databases in Ontario, Canada. Patients aged 40 years or older who visited a vascular surgeon between Apr. 1, 2004, and Mar. 31, 2007 (index date), and carried a diagnosis of PAD comprised the study cohort. The primary outcome was a composite of death or hospital admission for stroke or myocardial infarction. Secondary outcomes included lower limb amputation or revascularization. We used Cox proportional hazards modelling to compute unadjusted hazard ratios (HRs) and HRs adjusted for baseline covariates.

Results: A total of 6915 patients were studied, of whom 2461 (35.6%) were women. No significant differences in the risk of the primary outcome were observed between men and women (adjusted HR 0.99 [95% confidence interval (CI) 0.92–1.05]). Women were less likely than men to undergo minor amputation (adjusted HR 0.73 [95% CI 0.62–0.85]) and arterial bypass surgery (adjusted HR 0.82 [95% CI 0.71–0.94]) but were more likely to be admitted to hospital for acute myocardial infarction (adjusted HR 1.15 [95% CI 1.00–1.31]). There were no sex differences in the rates of major amputation or transluminal percutaneous angioplasty.

Interpretation: We identified no significant differences in the composite risk of major adverse cardiovascular events between women and men with PAD, although our findings suggest men may be at increased risk for adverse limb events compared with women. Cardiovascular health campaigns should focus on both women and men to promote early diagnosis and management of PAD.

eripheral arterial disease (PAD) is the third most common cause of death from cardiovascular disorders, after coronary heart disease and stroke. There has been a rapid increase in the number of people with PAD over the last decade, with about 202 million people with the disease globally in 2010.¹ Data from the Global Burden of Disease projects in 1990 and 2010 showed that, although men have an overall greater burden of PAD, women experienced a much more dramatic increase in PAD-related death and disability over that period than did men.² Despite this, PAD is not currently a focus in national cardiovascular health programs for men and women in Canada.³

Sex-based differences in the clinical presentation and outcomes of men and women with other vascular diseases such as coronary heart disease are well documented,⁴⁻⁶ and several national initiatives have been established to tackle these differences in health care. Differences between women and men in the etiology and clinical presentation of PAD have also been suggested.⁷ However, long-term data on sex-based differences

in the cardiovascular and limb outcomes of patients with PAD are limited. Previous population-based studies have focused primarily on postoperative in-hospital outcomes, and their findings have been conflicting.^{8–10} In a Scientific Statement published in 2012, the American Heart Association noted that "cardiovascular mortality, all-cause mortality, and major coronary event rates by gender have not been well defined in population-based studies."¹¹

To address this gap in knowledge, we conducted a large population-based cohort study to examine differences in rates of major adverse cardiovascular and limb events between

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women and men with PAD over the long-term. We hypothesized that women with PAD may be at increased risk for adverse cardiovascular events compared with men even after any differences in baseline characteristics are controlled for.

Methods

Setting

We designed a retrospective population-based cohort study of men and women with PAD using linked health care administrative databases in Ontario from Apr. 1, 2004, to Mar. 31, 2013. Over 13 million Ontario residents have access to universal health care services funded by the Ministry of Health and Long-Term Care.

Data sources

We used the following linked Ontario administrative data sets to conduct this study: the Canadian Institute for Health Information Discharge Abstract Database (captures all hospital admission information, including diagnoses and procedures), the Ontario Health Insurance Plan Database (data from all inpatient and outpatient physician billing claims), the Registered Persons Database (demographic and vital statistics data), the National Ambulatory Care Reporting System (data from outpatient surgical procedures and emergency department visits), the Institute for Clinical Evaluative Sciences Physicians Database (demographic and specialty information about all physicians practising in Ontario) and the Ontario Drug Benefit database (data from medication prescriptions dispensed to Ontario residents aged 65 years or older). These databases have shown to be of high quality,12 have been validated for identifying a variety of diagnoses and procedures¹³⁻¹⁶ and are routinely used for population-based outcome studies. 17-19

Patients

All men and women who were seen by a vascular surgeon in Ontario between Apr. 1, 2004, and Mar. 31, 2007, and who had a diagnosis of PAD within 3 years before that visit were considered eligible for the study. We defined the diagnosis of PAD using a validated model-based billing code algorithm that has been shown to be accurate in identifying patients with PAD in the noninvasive vascular laboratory and the community. Details of the billing code algorithm and the codes of the International Classification of Diseases, 9th and 10th editions, used to identify patients with PAD can be found in Appendix 1, available at www.cmajopen.ca/content/4/1/E124/suppl/DC1. To restrict the analysis to patients with atherosclerotic disease, we excluded patients who were less than 40 years old.

Outcome measures and follow-up

The primary outcome was a composite of death or hospital admission for acute myocardial infarction or stroke. Each individual component outcome of the primary outcome was examined as a secondary outcome. We also examined the following limb outcomes as secondary outcomes: major amputation (level of ankle, below-knee or above-knee), minor ampu-

tation (toe or through-foot), peripheral arterial bypass surgery and percutaneous transluminal angioplasty. Hospital admission for congestive heart failure, carotid endarterectomy, abdominal aortic aneurysm repair, coronary revascularization or initiation of dialysis was measured as a tertiary outcome. Follow-up was started on the day the patient had a clinical encounter with a vascular surgeon (cohort entry date). Patients were followed until they experienced the primary outcome or reached the maximum follow-up period of 7 years, or until the end of the study (Mar. 31, 2013).

Covariates

We controlled for confounding factors that might influence the effect of sex on adverse cardiovascular outcomes by using various linked administrative data sets. These factors included patient age, income level, medical comorbidities and overall comorbidity burden, medication use and use of health care services (including hospital admissions, emergency department visits and outpatient physician visits). We estimated income level by linking the patient's home postal code to population census data from Statistics Canada to obtain the median household income quintile.²¹ We estimated overall comorbidity burden using the Charlson Comorbidity Index with a 1-year look-back window in the data sets. This comorbidity index has been validated to predict the probability of death from comorbid disease in patients with PAD and aortic disease.²² In general, the risk of death from an increase of 1 Charlson Comorbidity Index score is equivalent to the risk of death from an additional decade of age.²² In addition, we controlled for patient ambulatory care grouping using the Comprehensive Ambulatory Classification System. This national classification system places patient visits into groups that are homogeneous both clinically and in terms of resource use. It can be used to further predict patient comorbidity burden and use of health care resources.²³ We controlled for baseline medication use by determining the medication prescriptions dispensed to patients using the Ontario Drug Benefit database.12 Details of codes used to define covariates and outcomes are outlined in Appendix 2, available at www.cmajopen.ca/content/4/1/E124/suppl/DC1.

Statistical analysis

We used standardized differences to compare male and female patients' baseline demographic and clinical characteristics. Standardized differences of less than 0.1 generally indicate good balance.²⁴

We performed time-to-event analyses using multivariable Cox proportional hazards regression (PROC PHREG in SAS, version 9.4) to compare outcomes of men and women, with men as the reference group. We generated unadjusted hazard ratios (HRs) and HRs adjusted for baseline demographic and clinical variables. We also calculated adjusted event rates per 1000 person-years for each outcome. We tested the proportional hazards assumption by visually inspecting log-log survival curves and by examining the statistical significance of time-dependent covariates. We also generated adjusted Kaplan–Meier curves with 95% confidence intervals (CIs) to



Characteristic	Men $(n = 4454)$	Women ($n = 2461$)	Standardized difference*	
Age, yr, mean ± SD	68.7 ± 10.6	71.6 ± 11.5		
Charlson Comorbidity Index score, no. (%)†				
0	219 (4.9)	133 (5.4)	0.02	
1	953 (21.4)	619 (25.2)	0.09	
≥2	2735 (61.4)	1384 (56.2)	0.11	
Neighbourhood income quintile, no. (%)†				
1 (lowest)	1149 (25.8)	641 (26.0)	0.01	
2	986 (22.1)	593 (24.1)	0.05	
3	833 (18.7)	432 (17.6)	0.03	
4	759 (17.0)	423 (17.2)	0.00	
5 (highest)	705 (15.8)	353 (14.3)	0.04	
Ambulatory care group, no. (%)†				
Nervous system diseases	122 (2.7)	61 (2.5)	0.02	
Eye diseases	7 (0.2)	≤ 5‡	0.01	
Ear, nose, mouth and throat diseases	34 (0.8)	23 (0.9)	0.02	
Respiratory diseases	113 (2.5)	77 (3.1)	0.04	
Circulatory system diseases	885 (19.9)	525 (21.3)	0.04	
Digestive system diseases	135 (3.0)	108 (4.4)	0.07	
Hepatobiliary and pancreatic diseases	24 (0.5)	6 (0.2)	0.04	
Musculoskeletal and connective tissue diseases	181 (4.1)	117 (4.8)	0.03	
Skin, subcutaneous tissue and breast diseases	418 (9.4)	239 (9.7)	0.01	
Endocrine, nutritional and metabolic diseases	325 (7.3)	126 (5.1)	0.09	
Kidney and genitourinary tract diseases	297 (6.7)	148 (6.0)	0.03	
Hematological diseases	26 (0.6)	17 (0.7)	0.01	
Oncological diseases	37 (0.8)	7 (0.3)	0.07	
Systematic infections	34 (0.8)	8 (0.3)	0.06	
Mental diseases	18 (0.4)	13 (0.5)	0.02	
Examination and other health factors	384 (8.6)	191 (7.8)	0.03	
Trauma, coma and toxic effects	258 (5.8)	147 (6.0)	0.01	
Health service use				
No. of outpatient physician visits in past yr, mean \pm SD	17.0 ± 11.1	16.6 ± 10.2	0.04	
No. of emergency department visits in past 3 yr, mean ± SD	3.9 ± 5.3	4.0 ± 5.1	0.01	
No. of hospital admissions in past 3 yr, mean ± SD	2.8 ± 2.3	2.8 ± 2.2	0.01	
Medications, no. (%)§	n = 2810	n = 1742		
Orally administered antihyperglycemic	798 (28.4)	382 (21.9)	0.15	
Insulin	505 (18.0)	263 (15.1)	0.08	
Statin	1715 (61.0)	949 (54.5)	0.13	
β-Blocker	1403 (49.9)	838 (48.1)	0.04	
Calcium-channel blocker	1086 (38.6)	831 (47.7)	0.18	
ACE inhibitor or ARB	1999 (71.1)	1206 (69.2)	0.04	
Diuretic	1017 (36.2)	701 (40.2)	0.08	
Any antihypertensive	2470 (87.9)	1545 (88.7)	0.02	
Clopidogrel	348 (12.4)	198 (11.4)	0.03	
Acetylsalicylic acid	604 (21.5)	365 (21.0)	0.01	
Warfarin	719 (25.6)	372 (21.4)	0.10	
Comorbid conditions, no. (%)				
Congestive heart failure	721 (16.2)	388 (15.8)	0.01	
Stroke	180 (4.0)	98 (4.0)	0.00	
Acute myocardial infarction	564 (12.7)	250 (10.2)	0.08	
Coronary artery disease	1173 (26.3)	495 (20.1)	0.15	
Chronic obstructive pulmonary disease	359 (8.1)	240 (9.8)	0.06	
Chronic kidney disease	871 (19.6)	403 (16.4)	0.08	
Diabetes mellitus	1694 (38.0)	776 (31.5)	0.14	

Note: ACE = angiotensin-converting enzyme, ARB = angiotensin-receptor blocker.

*Standardized difference of > 0.1 indicates significant difference.

†Missing values: Charlson Comorbidity Index 547 men (12.3%), 325 women (13.2%); neighbourhood income quintile 22 men (0.5%), 19 women (0.8%); ambulatory care group 1156 men (26.0%), 643 women (26.1%).

‡Cells containing ≤ 5 observations are suppressed.

§Medication use is reported only for those aged 66 years or older.



demonstrate the cumulative incidence of primary and secondary outcomes.

All statistical analyses were conducted with the use of SAS, version 9.4.

Ethical approval for this study was obtained from the Institute for Clinical Evaluative Sciences, and the need for individual patient consent was waived.

Results

A total of 6915 patients with PAD were included in our analysis. Of these patients, 2461 (35.6%) were women. The median duration of follow-up for the cohort was 5.4 years (interquartile range 1.8–7 years).

Baseline characteristics

Women with PAD were more likely than their male counterparts to be older (72 v. 69 years) (Table 1). Men were more likely than women to have a history of coronary artery disease or diabetes mellitus, and they tended to have a higher Charlson Comorbidity Index score, indicating a greater comorbidity burden. Men were also prescribed orally administered antihyperglycemic agents and statins at higher rates than women, whereas women were prescribed calcium-channel blockers at a higher rate than men. Analysis of baseline income quintile, use of health care services and clinical encounters based on 17 unique ambulatory care groupings showed no significant differences between women and men.

Primary outcome

The primary outcome (composite of death or hospital admission for acute myocardial infarction or stroke) occurred in 2860 men (64.2%) and 1644 women (66.8%). Women experienced the primary outcome at a higher rate in the unadjusted model (unadjusted HR 1.06 [95% CI 1.00 –1.13]), but this difference disappeared after baseline covariates were controlled for (adjusted HR 0.99 [95% CI 0.92–1.05]) (Table 2). The cumulative incidence of the primary outcome over follow-up is presented in Figure 1.

Secondary and tertiary outcomes

Women were less likely than men to undergo minor amputation (adjusted HR 0.73 [95% CI 0.62–0.85]) and arterial bypass surgery (adjusted HR 0.82 [95% CI 0.71–0.94]) but were more likely to be admitted to hospital for acute myocardial infarction (adjusted HR 1.15 [95% CI 1.00–1.31]) (Table 2). There were no differences in the rates of major amputation, percutaneous transluminal angioplasty, stroke or death between women and men. The cumulative incidences of secondary limb outcomes are presented in Figures 2 and 3. In addition, there were no significant differences in any of the tertiary outcomes between men and women in the adjusted models.

Interpretation

In this large population-based cohort study, we found no statistically significant differences between women and men with

Table 2: Outcomes by sex									
	Men,	Women,	Lipodiusto d LID	Adjusted LID*		Adjusted event rate per 1000 person-years			
Outcome	no. (%) (n = 4454)	no. (%) (n = 2461)	Unadjusted HR (95% CI)	Adjusted HR* (95% CI)	Men	Women			
Primary outcome†	2860 (64.2)	1644 (66.8)	1.06 (1.00–1.13)	0.99 (0.92–1.05)	162.6	159.4			
Secondary outcome									
Acute myocardial infarction	618 (13.9)	360 (14.6)	1.07 (0.94–1.22)	1.15 (1.00–1.31)	30.1	34.5			
Stroke	229 (5.1)	140 (5.7)	1.13 (0.92–1.40)	1.02 (0.82–1.27)	9.6	9.7			
Death	2562 (57.5)	1468 (59.6)	1.06 (0.99–1.13)	0.96 (0.90-1.03)	130.0	124.1			
Major amputation	764 (17.2)	375 (15.2)	0.89 (0.79–1.01)	0.95 (0.84–1.08)	38.3	36.2			
Minor amputation	583 (13.1)	211 (8.6)	0.65 (0.55–0.76)	0.73 (0.62–0.85)	23.8	16.9			
Arterial bypass surgery	685 (15.4)	311 (12.6)	0.82 (0.72-0.94)	0.82 (0.71–0.94)	39.2	31.5			
PCTA	609 (13.7)	330 (13.4)	0.99 (0.87–1.14)	1.04 (0.90–1.19)	33.0	34.1			
Tertiary outcome									
Congestive heart failure	1144 (25.7)	623 (25.3)	1.01 (0.92–1.11)	1.01 (0.91–1.12)	60.9	61.4			
Carotid endarterectomy	82 (1.8)	34 (1.4)	0.76 (0.51–1.14)	0.77 (0.50–1.12)	0.4	0.3			
Abdominal aortic aneurysm repair	271 (6.1)	126 (5.1)	0.85 (0.69–1.05)	0.92 (0.74–1.15)	9.1	8.3			
Coronary revascularization	442 (9.9)	187 (7.6)	0.77 (0.65–0.91)	0.85 (0.71–1.02)	18.2	15.5			
Initiation of dialysis	749 (16.8)	322 (13.1)	0.77 (0.68–0.88)	0.89 (0.78–1.03)	21.4	18.2			

Note: CI = confidence interval, HR = hazard ratio, PCTA = percutaneous transluminal angioplasty.

†A composite of death or hospital admission for acute myocardial infarction or stroke

^{*}Adjusted for patient age, income level, medical comorbidities and overall comorbidity burden, medication use, use of health care services (including hospital admissions, emergency department visits and outpatient physician visits) and ambulatory care group classification.

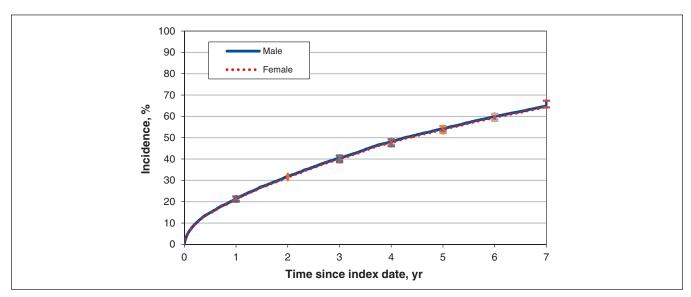


Figure 1: Incidence over follow-up of primary outcome (death or hospital admission for acute myocardial infarction or stroke), adjusted for patient age, income level, medical comorbidities and overall comorbidity burden, medication use, use of health care services (including hospital admissions, emergency department visits and outpatient physician visits) and ambulatory care group classification. Error bars represent 95% confidence intervals.

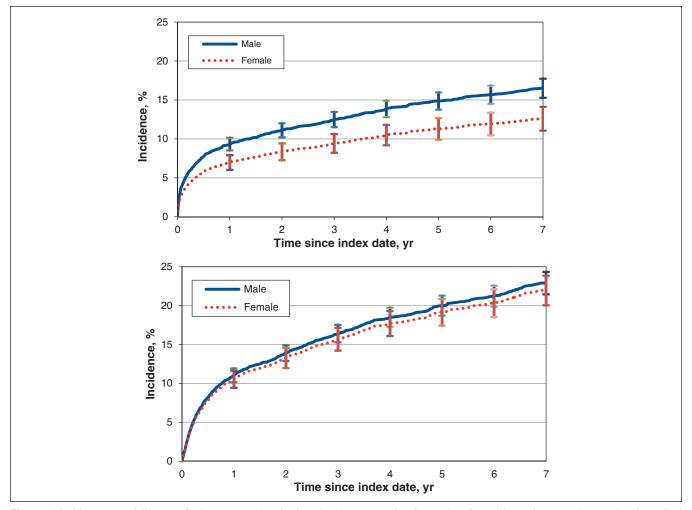


Figure 2: Incidence over follow-up of minor amputation (top) and major amputation (bottom), adjusted for patient age, income level, medical comorbidities and overall comorbidity burden, medication use, use of health care services (including hospital admissions, emergency department visits and outpatient physician visits) and ambulatory care group classification. Error bars represent 95% confidence intervals.

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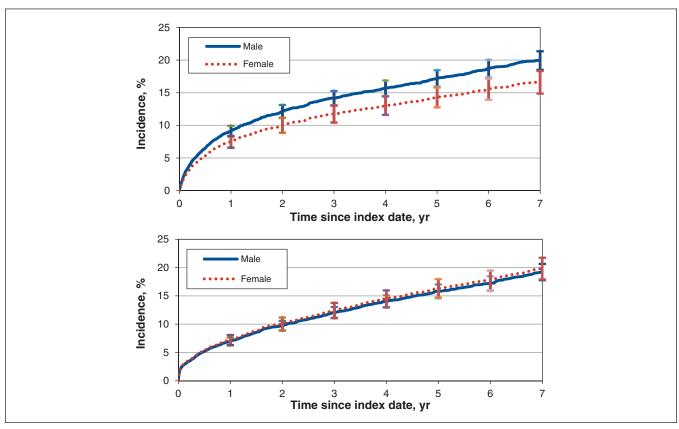


Figure 3: Incidence over follow-up of arterial bypass surgery (top) and percutaneous transluminal angioplasty (bottom), adjusted for patient age, income level, medical comorbidities and overall comorbidity burden, medication use, use of health care services (including hospital admissions, emergency department visits and outpatient physician visits) and ambulatory care group classification. Error bars represent 95% confidence intervals.

respect to a composite risk of acute myocardial infarction, stroke or death. Secondary analyses showed that female sex was associated with a higher risk of hospital admission due to myocardial infarction but a lower risk of minor amputation and arterial bypass surgery.

The findings of prior studies examining sex-related differences in the outcomes of patients with PAD following peripheral arterial procedures have been conflicting. Jackson and colleagues8 recently reported data for 12 379 patients (41% female) who underwent peripheral vascular interventions. Although women experienced more procedure-related complications, the rates of myocardial infarction, stroke, death and amputation did not differ between women and men. However, the study was limited to in-hospital postprocedure follow-up. In contrast, Vouyouka and colleagues9 analyzed population-based data from 372 692 PAD-related surgical hospital admissions (44% involving women) and reported a higher risk of postoperative complications, major amputation and death among women even after adjusting for baseline covariates. That study was also limited by a lack of longitudinal follow-up after discharge, and only patients undergoing peripheral arterial revascularization or amputation procedures were enrolled. Recent data from the US Nationwide Inpatient Sample suggest that women who undergo interventions for PAD have higher rates of in-hospital death than do men, although the rates for both sexes have declined over time,

likely owing to improvements in technology and medical therapy.¹⁰ The authors measured no other clinical outcomes owing to a limitation of the administrative databases.

Our study has several strengths that provide insight into the influence of sex on PAD outcomes. First, unlike prior studies, the index event in our study cohort was not a surgical intervention for PAD but, rather, a clinical encounter with a vascular surgeon and a prior diagnosis of PAD. This approach allowed us to include inpatients and outpatients with varying levels of disease severity in our cohort. Prior studies focused on patients with advanced PAD (critical limb ischemia or debilitating intermittent claudication) undergoing peripheral arterial interventions.^{8-10,25,26} Findings from our study may be more generalizable because we included a broader range of patients. Second, our study provides robust long-term data (median follow-up duration 5.4 years), whereas prior population-based studies focused largely on in-hospital outcomes in this population. Third, we adjusted for several factors that may confound the relationship between sex and PAD outcomes, such as socioeconomic status, use of health care services, ambulatory care grouping and overall comorbidity burden. These factors play an important role in access to care and have not been measured in prior studies examining sex-related differences in PAD. Finally, our study provides insight into characteristics of patients with PAD that have not been previously described. For instance, we found that about 50% of

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men and women with PAD belonged to the lowest 2 income quintiles. Given that socioeconomic class is closely associated with health status,²⁷ our data indicate further investigation is warranted to determine the impact of lower socioeconomic status on the health outcomes of patients with PAD.

In our cohort, women with PAD tended to be older than men with the disease, which is consistent with data from other studies.8-10,25 This may be because women are more likely to experience atypical symptoms that are not characteristic of a classical history of intermittent claudication, thus leading to a delay in PAD diagnosis.^{28,29} Furthermore, women with PAD had more advanced disease at the time of presentation, as evidenced by a greater incidence of critical limb ischemia^{30,31} and faster functional decline.32 Although these clinical findings can be explained by advanced age, other factors, such as an enhanced proinflammatory state, 33 impaired microvascular function³⁴ and greater burden of femoropopliteal lesions,³⁵ observed in women with PAD may also contribute to more advanced disease compared with that in men.

Although women were at increased risk for acute myocardial infarction, stroke or death in the unadjusted analysis, we found no differences among women and men after adjusting for age and other covariates. This further indicates the importance of early diagnosis and treatment of PAD in women. We did, however, find 2 key differences between men and women in our secondary analyses that merit discussion. First, we found that women were less likely than men to undergo minor amputation and arterial bypass surgery. This may be because men were observed to have a greater comorbidity burden at baseline, including a prior history of diabetes mellitus, which is a significant risk factor for advanced atherosclerosis and amputation.¹⁹ Second, we found that women were at increased risk for acute myocardial infarction in the adjusted analysis. Given that women in our study were on average 3 years older than men, this finding is consistent with previous research indicating that women experience acute myocardial infarction at an older age than men.³⁶ However, the possibility also exists that a lower rate of peripheral arterial interventions signifies that women with PAD may be inappropriately undertreated with minor amputation and bypass surgery compared with men; we were not able to confirm this owing to a lack of data about severity of PAD in our databases. We found no significant differences between women and men in the adjusted rates of stroke, death, major amputation, percutaneous transluminal angioplasty or any of the tertiary outcomes.

Limitations

Our study has several limitations. First, as with all observational studies, our results may be biased by unmeasured differences between men and women, such as body mass index, smoking history, blood pressure, blood lipid values and exercise level. However, we measured several demographic and clinical covariates at baseline, and we also adjusted our analyses for variables that may confound the relationship between sex and PAD outcomes. Second, inaccurate coding is a possible source of bias in population-based studies. To mitigate this risk, we used codes that had already been validated in our databases to identify patients with PAD and their outcomes. The model-based billing code algorithm has good accuracy (86% sensitivity, 83% specificity, 91% positive predictive value, 74% negative predictive value) in identifying patients with PAD.²⁰ Coding for the major cardiovascular and limb outcomes of interest, including acute myocardial infarction (83% sensitivity, 87% positive predictive value), stroke (81%) accuracy), arterial bypass surgery (87% sensitivity, 88% positive predictive value) and major amputation (98% sensitivity, 96% positive predictive value), has also been validated in our databases through reabstraction studies.37 Finally, our databases did not allow us to determine the severity of PAD symptoms in our cohort. However, by including patients who had a clinical encounter with a vascular surgeon, we are confident that our cohort represents a broad PAD population with varying severity of disease that is managed by vascular specialists.

Conclusion

Our results suggest that, although women with PAD tend to be older than their male counterparts, no sex-based differences exist in the long-term composite risk of acute myocardial infarction, stroke or death in patients with PAD after baseline risk factors are controlled for. Men with PAD tend to have a higher comorbidity burden than women with the disease, and they may be at increased risk for adverse limb events. Our results highlight the need to target both women and men in cardiovascular health campaigns to promote early diagnosis and management of PAD. Formal research into the development and implementation of such programs is warranted.

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